Basics of regression analysis

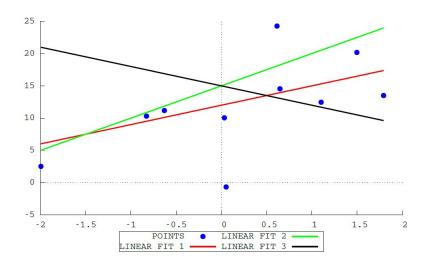
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Regression analysis

• Assume we have a collection of data on two economic quantities x and y for n individuals or unit of analysis, that is: $\{(x_i, y_i; i = 1, ..., n)\}$

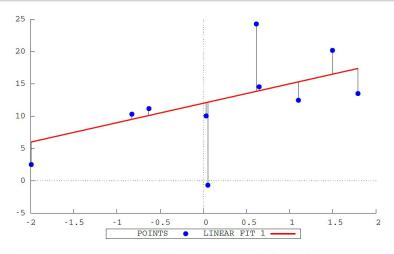
- Suppose further that we would like to describe the relation between *Y* and *X* through the linear relation: $y = \beta_0 + \beta_1 x$
- ⇒ How do we get the values of the parameters?



- Each line corresponds to a different slope and a different intercept
 - Linear Fit 1: $\hat{\mathbf{y}} = 12 + 3 \times x$
 - Linear Fit 2: $\hat{y} = 15 + 5 \times x$
 - Linear Fit 1: $\hat{y} = 15 3 \times x$
- We need a way to assess which line is best description of the data
- A possible criterion to decide the best one is to start from the error we make by using each line instead of the data points.

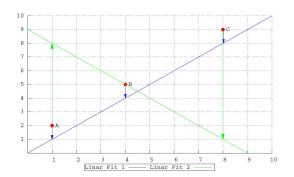
$$u_i = \hat{y}_i - y_i \tag{1}$$





Black segments represent the error we make by substituting y_i with \hat{y}_i , that is $error_i = y_i - \hat{y}_i = y_i - \beta_0 - \beta_1 x_i$. In this example is $y_i - 12 - 3x_i$.

- For each pair (x_i, y_i) you have an error u_i , so we can compute an *overall* error just by summing, that is $\sum_i u_i$
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 - ✓ Linear Fit 1: errors are (1,1,1), so Sum of errors=3
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 To balance-out the effect of cancellations between positive and negative errors, we define the Sum of Squared Residuals

$$SSR = \sum_{i=1}^{n} \hat{u}_{i}^{2} = \sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}$$

- We take this as the overall size of the mistake made when using each linear fit
- \implies A natural criterion to estimate the parameter is to find values of β_0 and β_1 that minimize the SSR: this is called Ordinary Least Square (OLS) method

The OLS criterion for Linear Fitting

Formally the problem is

$$\underset{\hat{\beta}_0, \hat{\beta}_1}{\text{MIN}} \quad \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

The two necessary conditions to identify the solution read

$$\begin{cases} \frac{\partial \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2}{\partial \hat{\beta}_0} = 0 \\ \frac{\partial \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2}{\partial \hat{\beta}_1} = 0 \end{cases} \begin{cases} -2 \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \\ -2 \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) x_i = 0 \end{cases}$$

with solutions

$$\begin{cases} & \hat{\beta}_{0} = \bar{Y} - \hat{\beta}_{1}\bar{X} \\ & \hat{\beta}_{1} = \frac{\sum_{i}(x_{i} - \bar{x})(y_{i} - \bar{y})}{\sum_{i}(x_{i} - \bar{x})^{2}} = \frac{COV(x, y)}{VAR(x)} \end{cases}.$$



The OLS criterion for Linear Fitting

• Once we have the numbers $\hat{\beta}_0$ and $\hat{\beta}_1$ for a given data set, we write the OLS fitted line as a function of x:

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

- The OLS fitted line allows us to predict y for any (sensible) value of x.
- The intercept, $\hat{\beta}_0$, is the predicted y when x = 0. (The prediction is usually meaningless if x = 0 is not possible.)
- The slope, $\hat{\beta}_1$, allows us to predict changes in y for any (reasonable) change in x:

$$\Delta \hat{y} = \hat{\beta}_1 \Delta x$$

■ If $\Delta x = 1$, so that x increases by one unit, then $\Delta \hat{y} = \hat{\beta}_1$.



Algebraic properties of OLS

$$\underset{\hat{\beta}_0, \hat{\beta}_1}{\text{MIN}} \quad \sum_{i} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

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$$\begin{cases} \sum_{i} \hat{u}_{i} = 0 \end{cases} \qquad \begin{bmatrix} \text{AP1. OLS residuals always add up to} \\ 0 \end{cases}$$
$$\sum_{i} \hat{u}_{i} x_{i} = 0 \end{cases} \qquad \begin{bmatrix} \text{AP2. Sample covariance between residuals } \hat{u}_{i} \text{ and } x_{i} \text{ is } 0. \end{cases}$$

- Suppose to have observations on a population of individuals (even if most often we work with samples)
- We have data (Y_i, X_i) , where this time X is a set of characteristics (not just one variable)
- We are interested in understanding to what extent knowledge of Xs helps to characterize Y, or, similarly, to explain or predict Y on the basis of the Xs. That is, we are interested in some function of Y conditional on the Xs
 - Y is the dependent variable
 - The Xs are called covariates (aka regressors or explanatory variables)
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Theorem (CEF decomposition properties)

$$Y_i = E[y_i|X_i] + \epsilon_i ,$$

where

- $\blacksquare E[\epsilon_i|X_i] = 0;$
- \bullet ϵ_i is uncorrelated with any function of X_i .
- \implies Any variable y_i can be decomposed into two pieces:
 - A piece which is orthogonal to any function of the Xs
 - A piece that is explained by the Xs, captured by the CEF
 - $E[Y_i \mid X_i] = \int dt \ t \ f_y(t \mid X_i = x)$, with f_y a conditional density, is called the Conditional Expectation Function (CEF)



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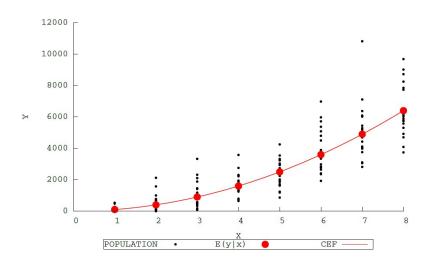
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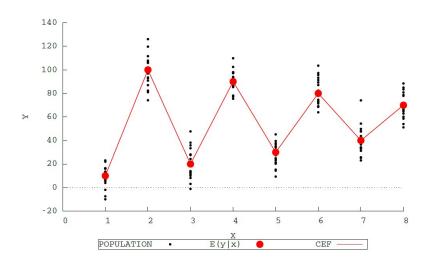
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CEF example 1



CEF example 2



Theorem (CEF prediction property)

Let $m(X_i)$ be any function of X_i . Then

$$E[y_i|X_i] = \underset{m(X_i)}{arg \ min} \ E\left[(y_i - m(X_i))^2\right]$$

so the CEF is the MMSE predictor of y_i given X_i .

- ⇒ The CEF is the best predictor of *Y* given the Xs, among *all possible* functions of the Xs (in MSE terms)
 - ✓ The CEF represents a "precise" way to characterize the relationship between Y and the Xs, if we ask "how can we explain or predict Y, based on info about the Xs?"

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Why we like OLS? The CEF-OLS link

Theorem (the regression-CEF theorem)

The function $X_i'\beta$ provides the MMSE linear approximation to $E[y_i|X_i]$, that is

$$\beta = \underset{b}{arg min} E\left\{ (E[y_i|X_i] - X_i'b)^2 \right\} . \tag{2}$$

Proof.

Write

$$(y_i - X_i'b)^2 = (y_i - E[y_i|X_i])^2 + (E[y_i|X_i] - X_i'b)^2 + 2(y_i - E[y_i|X_i])(E[y_i|X_i] - X_i'b) .$$

The first term does not involve b and the last one has expectation zero by the CEF-decomposition property.

- \implies Problem (2) has same solution as MIN_b $E\left[(Y_i X_i'b)^2\right]$, which is exactly what we do in OLS !!!
 - ✓ The linear regression function X'b that we get from OLS is a good approximation (the best in MMSE terms) of the CEF

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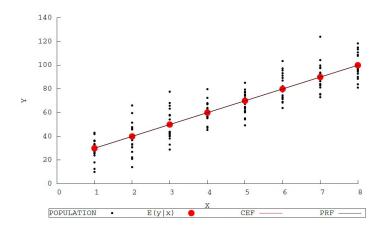
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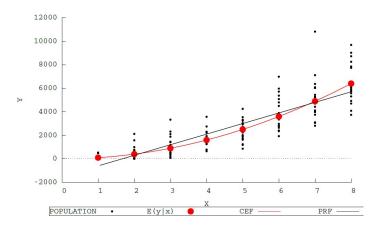
How good is the OLS-Population Regression Function?

If the CEF is linear, then the (linear) Population Regression Function is exactly the CEF



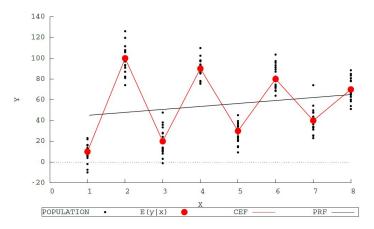
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In other situations, it is the best we can do (in MMSE terms), but not always satisfactory



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• We have learnt that each y_i can be expressed as

$$y_i = E(y|x_i) + u_i ,$$

where the error term u_i captures how much we are distant from the CEF

- What is in the error term?
 - omitted factors, due to a wrong idea or theory about what we should consider as predictors (or determinants) of y
 - omitted factors due to lacking data on some Xs that we would like to include
 - wrong functional form



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• Crucial property is the Zero Conditional Mean (ZCM) property:

ZCM. The average, or expected, value of u, conditional on x, is zero. Formally, E(u|x)=0,

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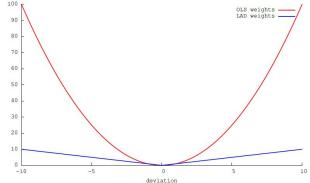
- Do not forget that OLS confine the attention to the expected value of the conditional distribution of *y* give *X*
 - In general, one may be interested into other features of the same distribution, so we would need different techniques in that case
 - The conditional expected value might be particularly meaningless if *y* has a very skewed distribution, since in that case the mean of *y* says little
- Do not forget that the OLS weights a lot large errors (taking squares of the distance from linear fit line)
 - This means that few outliers can dramatically influence the estimates of the parameters
 - Often you either drop the outliers (if they are just few data-points) or look for different techniques that are less influenced by outliers



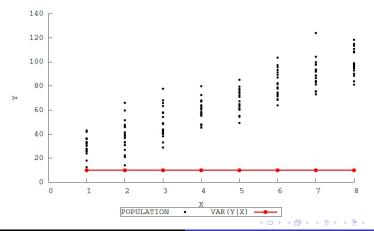
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 One popular and easy correction for outliers is Least Absolute Deviation (LAD)

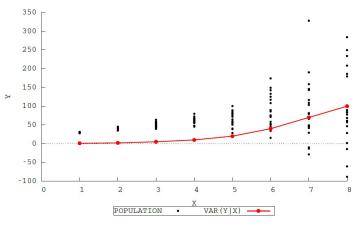
$$\underset{\hat{\beta}_{0},\hat{\beta}_{1}}{\text{MIN}} \sum_{i} (y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}x_{i})^{2} \qquad \underset{\hat{\beta}_{0},\hat{\beta}_{1}}{\text{MIN}} \sum_{i} \left| y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}x_{i} \right|$$
OLS estimation
$$\text{LAD estimation}$$



• Another property worth noticing is HOMOSKEDASTICITY, meaning that the variance of the error (and thus of the part of Y not captured by the regression) is the same for every value of x: $VAR(u \mid x) = \sigma^2 > 0$, $\forall x$



• HETEROSKEDASTICITY, instead, means that the variance of u (or of the part of Y not captured by the regression line) varies with the values of x



Regression analysis: final remarks

- This was a highly simplified presentation: many other problems remain in practical work
- The practice of econometrics is mostly to deal with real-world situation where OLS assumptions are difficult to maintain (e.g., recall sample-selection or endogeneity discussed in Gibrat's regression)
- Maximum-likelihood is a general alternative, flexible and able to also account for non-linearities: the idea is to assume a distribution for the errors, and then write the joint probability density and maximize it (e.g., recall above discussion about parametric density estimation, or the non-linear estimation of the scaling relationship between variance of growth and size)
- Nevertheless, it is useful as a benchmark: economists like to frame their research questions as "Is there an impact of a certain variable X on the value of an outcome value Y, and if so how strong is it?"